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ABSTRACT

This paper warns about the incorrect use of the popular Jarque-Bera test for normality of residuals in the case of small and medium-size samples. It also provides a natural modification of the test that mitigates the problem.

Keywords: Normality test; Omnibus test

JEL classification: C10, C20

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1. INTRODUCTION

The test for univariate normality of observations and residuals introduced by Jarque and Bera (1980, 1987) has gained great acceptance among economists. It is an omnibus test based on the standardized third and fourth moments:

$$LM = n[(\sqrt{b_1})^2/6 + (b_2-3)^2/24] \quad (1)$$

where n is the number of observations, $\sqrt{b_1} = m_3/m_2^{3/2}$, $b_2 = m_4/m_2^2$, and m_i is the i -th central moment of the observations [i.e., $m_i = \Sigma(x_j - \bar{x})^i/n$]. Asymptotically, the hypothesis of normality is rejected at some significance level if the value of LM exceeds the critical value of a chi-squared with two degrees of freedom. In the more usual case of a regression, (1) is calculated using the estimated residuals.

As shown by Jarque and Bera (1987), the test performs quite well compared to others available in the literature. This is not surprising since they proved that, if the alternatives to the normal distribution are in the Pearson family, LM is the corresponding Lagrange multiplier test for normality. Urzúa (1989) also showed the same when the alternatives are the maximum-entropy ("most likely") distributions with finite moments defined in Urzúa (1988).

However, the good performance of the test is highly dependent on the use, through a Montecarlo simulation, of empirical significance points (something, by the way, that is almost never done in studies where (1) is used). This is so because of the slow convergence in distribution to the chi-squared.

Interestingly enough, (1) has been known among statisticians since the work of Bowman and Shenton (1975). They derived it after noting that, under normality, the asymptotic means of $\sqrt{b_1}$ and b_2 are 0 and 3, the asymptotic variances are $6/n$ and $24/n$, and the asymptotic covariance is zero. Thus, LM is just the sum of squares of two asymptotically independent standardized normals.

Yet, there are few (if any) instances in the statistics literature where the Bowman-Shenton-Jarque-Bera test has been used. As one author flatly states in a comprehensive survey of tests for normality: "Due to the slow convergence of b_2 to normality this test is not useful." (D'Agostino, 1986, p. 391).

2. A NEW, BETTER-BEHAVED TEST STATISTIC

This section presents a new, better-behaved omnibus test for normality that is a natural extension of the Jarque-Bera test. The idea is straightforward: instead of the asymptotic means and variances of the standardized third and fourth moments, use their exact means and variances. Under normality, the latter can be easily computed using results already known to Fisher (1930).

Fisher's results were stated in terms of the so-called k -statistics, which can be expressed in terms of moments as (see Stuart and Ord, 1987, pp. 392, 422):

$$k_2 = nm_2 / (n-1), \quad k_3 = n^2 m_3 / (n-1)(n-2), \quad k_4 = n^2 [(n+1)m_4 - 3(n-1)m_2^2] / (n-1)(n-2)(n-3)$$

For our purposes, his relevant derivations are that, under normality, k_2 is independent of $k_p/k_2^{p/2}$ for $p=3,4,\dots$, and that

$$\text{var}(k_3/k_2^{3/2}) = 6n(n-1)/(n-2)(n+1)(n+3), \quad \text{var}(k_4/k_2^2) = 24n(n-1)^2/(n-3)(n-2)(n+3)(n+5)$$

But then we can use those results to easily show that, under normality, the exact mean and variance of the standardized third and fourth moments are

$$E(\sqrt{b_1}) = 0, \quad \text{var}(\sqrt{b_1}) = 6(n-2)/(n+1)(n+3) \quad (2)$$

$$E(b_2) = 3(n-1)/(n+1), \quad \text{var}(b_2) = 24n(n-2)(n-3)/(n+1)^2(n+3)(n+5) \quad (3)$$

And hence, using (2) and (3), we can finally define the new test, to be called the *adjusted Lagrange multiplier* test for normality, as:

$$ALM = n[(\sqrt{b_1})^2/\text{var}(\sqrt{b_1}) + (b_2 - E(b_2))^2/\text{var}(b_2)] \quad (4)$$

This new test statistic converges to the chi-squared with two degrees of freedom faster than the Jarque-Bera statistic, as can be glimpsed from the Montecarlo simulations reported in Table 1 (a more complete table is available upon request). Incidentally, the estimated significance points in that table can be used to test for normality of observations, but they cannot be used in the case of regression residuals, since, for each particular regression, the

significance points depend on the design (regressor) matrix and the distribution of the residuals (see, e.g., Weisberg, 1980).

3. ESTIMATED POWER OF THE TESTS

This section compares the power of *ALM* and *LM* when used as tests for normality of regression residuals. The Montecarlo simulation procedures used by us were, on purpose, identical to the ones employed by White and MacDonald (1980) in their much quoted paper on the subject. As in there, the five alternatives to the normal distribution of the residuals were: Student's *t* with five degrees of freedom; heteroscedastic normal; chi-squared with two degrees of freedom; Laplace (double exponential); and lognormal (all of them standardized to have mean zero and variance 25). Furthermore, for the generation of pseudo-random numbers we followed in each case the same computational procedure as in their paper.

Also following White and MacDonald (1980), the design matrices for the regressions were constructed adding to a column of ones three columns of uniform random numbers with mean zero and variance 25. The number of rows in each design matrix (i.e., the sample size) was given by $n = 20, 35, 50, 100$.

As a first exercise, we estimated the power of both tests when, as is incorrectly done in almost all empirical studies, the significance point is taken to be $\chi^2_{2,0.10} = 4.61$, even though the sample sizes are not large. The number of replications in each Montecarlo simulation was 10000 (instead of 200 in White and MacDonald, 1980), and the results are presented in Table 2.

As can be appreciated there, the results are overwhelmingly in favor of the new *ALM* test. It comes first in all the distributions considered and all the sample sizes. Furthermore, in the case of the smaller samples the power of *ALM* is significantly larger than the power of *LM*.

Naturally, the next question to ask is whether the same relative performance is obtained when, prior to applying the tests, "correct" significance points are found for each test using Montecarlo simulations (this is actually the procedure that is explicitly suggested in Jarque and Bera, 1987). The results

obtained that way are presented in Table 3. Once again, *ALM* outperforms *LM*. Interestingly enough, the only five cases (out of 20) where the *LM* test comes first correspond to the distributions that are farther apart from the normal.

4. CONCLUDING REMARKS

This paper has presented a new omnibus test for normality of residuals and observations: the adjusted Lagrange multiplier test *ALM*. As shown here, the *ALM* test outperforms in terms of power the traditional Jarque-Bera *LM* test, both, when significance points are directly taken from a chi-squared, or when the "correct" significance points are obtained through simulations. Thus, the use of *ALM* over *LM* seems warranted in both circumstances.

As a final point, a similar adjustment to the one suggested here can be extended to the multivariate tests for normality that are also based on third and fourth standardized moments, such as the one proposed in Urzúa (1989).

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Table 1

Significance points for two tests for normality of observations

<i>n</i> :	20	50	100	200	400	800	∞
<i>ALM</i>							
$\alpha=.10$	3.95	4.00	4.12	4.30	4.39	4.47	4.61
$\alpha=.05$	7.01	6.60	6.29	6.17	6.04	5.97	5.99
<i>LM</i>							
$\alpha=.10$	2.13	2.90	3.14	3.48	3.76	4.32	4.61
$\alpha=.05$	3.26	4.26	4.29	4.43	4.74	5.46	5.99

Sources: For *LM* Jarque and Bera (1987, table 2), and for *ALM* own simulations using 10000 replications.

Table 2

Tests for normality of residuals; estimated power with 10000 replications, using as significance point $\chi^2_{2,0.10}$

<i>n</i>		<i>t</i> ₅	Heteroskedastic		Laplace	Lognormal
			Normal	χ^2_2		
20	ALM	0.231	0.091	0.493	0.290	0.808
	LM	0.140	0.039	0.380	0.181	0.727
35	ALM	0.362	0.116	0.829	0.464	0.985
	LM	0.293	0.077	0.782	0.374	0.978
50	ALM	0.467	0.128	0.963	0.595	1.000
	LM	0.406	0.093	0.950	0.513	0.999
100	ALM	0.694	0.161	1.000	0.835	1.000
	LM	0.658	0.135	1.000	0.797	1.000

Table 3

Tests for normality of residuals; estimated power with 10000 replications, using estimated significance points ($\alpha = .10$)

<i>n</i>		<i>t</i> ₅	Heteroskedastic	χ^2_2	Laplace	Lognormal
			Normal			
20	ALM	0.254	0.477	0.533	0.317	0.831
	LM	0.247	0.456	0.586	0.306	0.856
35	ALM	0.391	0.724	0.864	0.494	0.989
	LM	0.376	0.697	0.896	0.470	0.992
50	ALM	0.493	0.852	0.973	0.624	1.000
	LM	0.474	0.831	0.982	0.595	1.000
100	ALM	0.712	0.984	1.000	0.849	1.000
	LM	0.698	0.981	1.000	0.836	1.000

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